# Approximate Bayesian Computation with Domain Expert in the Loop

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The University of Manchester

# Outline

#### Introduction

- 2 Approximate Bayesian Computation (ABC)
- 3 Failure of regression-ABC methods
- 4 Human-in-the-loop ABC
- 6 Results & Conclusion

#### Inference from data

Setting:

- Let data  $\mathbf{y}_{\text{obs}} = \{y_{\text{obs},i}\}_{i=1}^{n}$  be denoted by empirical distribution  $\mathbb{Q}^{n}$ .
- Model M<sub>Θ</sub> = {ℙ<sub>θ</sub> : θ ∈ Θ ⊂ ℝ<sup>q</sup>} is a parametric family of distributions.

**Estimation problem:** Given data  $\mathbf{y}_{obs}$ , estimate  $\theta$  s.t.  $\mathbb{Q}^n$  is "closest" to  $\mathbb{P}_{\theta}$ .

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Classical estimation techniques such as

Bayesian inference:  $p(\theta|\mathbf{y}_{obs}) \propto p(\mathbf{y}_{obs}|\theta)p(\theta)$ Maximum Likelihood (ML):  $\hat{\theta}_{ML} = \underset{\theta}{\operatorname{argmax}} p(\mathbf{y}_{obs}|\theta)$ 

require access to the likelihood function.

#### Problem: Many models have intractable likelihoods

The likelihood function cannot be evaluated numerically, or approximated in reasonable computation time.

Therefore, standard estimation techniques are unrealizable.

#### Causes of intractable likelihood:

- The model is simply too complex.
- Variables that are important for model description are unobserved.
- The likelihood function has not been derived yet for a newly constructed model.

Such models are called:

- Simulators
- Implicit models
- Generative models

#### Simulators in the Sciences

Physical sciences and engineering:

- Population genetics [Pritchard et al., 1999]
- Ecology and evolution [Beaumont, 2010]
- Astrophysics [Akeret et al., 2015]
- Epidemiology [Kypraios et al., 2017]
- Radio communications [Bharti et al., 2021]
- Atmospheric science [Kopka et al., 2016]
- Economics [Dyer et al., 2022]

**Solution:** use likelihood-free inference methods based on simulating from the model

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# Approximate Bayesian Computation (ABC)

ABC is a likelihood-free inference method that permits sampling from the approximate posterior of a model, given that it is easy to simulate from.

Rejection ABC algorithm

- Sample  $\theta^* \sim p(\theta)$
- Simulate data from model,  $\mathbf{y}^* \sim \mathbb{P}_{\theta^*}$
- If  $\rho(S(\mathbf{y}_{obs}), S(\mathbf{y}^*)) < \epsilon$ , accept  $\theta^*$

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#### Here

- $\rho(\cdot, \cdot)$  is a distance metric (typical choice is Euclidean distance)
- $S(\cdot)$  is the summarizing function
- $\epsilon$  is a tolerance threshold
- Accepted samples  $\theta_1, \ldots, \theta_N$  are iid from the approximate posterior:

 $p( heta|
ho(\mathcal{S}(\mathbf{y}_{\mathrm{obs}}),\mathcal{S}(\mathbf{y})<\epsilon) pprox p( heta|\mathbf{y}_{\mathrm{obs}})$ 

The "approximation" in Bayesian inference arises due to

- use of tolerance threshold in accepting parameter samples
- summarizing the data into a few statistics. If  $S(\cdot)$  is a *sufficient statistic* of **y**, then

$$p( heta|
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Ingredients required for implementing an ABC algorithm:

- distance metric  $\rho(\cdot, \cdot)$
- summary statistics  $S(\cdot)$
- tolerance threshold  $\epsilon$

 $\rho(S(\mathbf{y}_{obs}), S(\mathbf{y}^*)) < \epsilon$ 

Choosing  $\epsilon$ 

• In practice, select  $\epsilon$  as a small percentile of the simulated distances, i.e. given  $\{(\theta_i^*, S(\mathbf{y}_i^*))\}_{i=1}^M$ , accept the  $\epsilon M$  samples of  $\theta_i^*$  with the least  $\rho(S(\mathbf{y}), S(\mathbf{y}_i^*))$ 

Choose  $\rho(\cdot, \cdot)$ 

- Euclidean distance for summary-based ABC
- Integral probability metrics such as maximum mean discrepancy, Wasserstein distance

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Choosing  $S(\cdot)$ 

- The choice of statistics is non-trivial as it involves trade-off between
  - information loss due to summarization
  - 2 curse of dimensionality
- Fundamental unsolved problem in ABC

### **Choosing Summary Statistics**

In practice, domain experts manually handcraft and select statistics:

- laborious and time-consuming
- involves multiple trial-and-error steps
- takes up majority of the time of likelihood-free inference projects

Hence, domain knowledge is vital for constructing summaries.

## **Choosing Summary Statistics**

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Hence, domain knowledge is vital for constructing summaries.

Existing methods given a pool of candidate statistics:

- Subset selection (Joyce & Marjoram, 2008)
- Projection techniques (Fearnhead & Prangle, 2012)
- Regression adjustment (Beaumont et al., 2010)

However, performance of these methods degrade when

- number of available model simulations is limited (low-simulation regime)
- model is misspecified, i.e.,  $\mathbb{Q}^n \notin \mathcal{M}_\Theta$

# Outline

#### Introduction

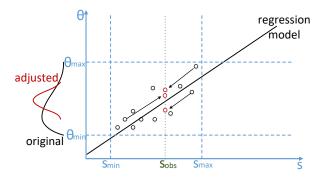
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#### Case study: Regression-ABC methods

Given accepted samples  $(\theta_i, \mathbf{s}_i)_{i=1}^{n_e}$ , regression-ABC methods account for the difference between the simulated and observed statistic by adjusting the parameter values using model

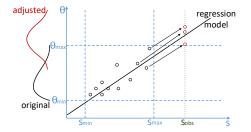
$$\theta_i = \varphi(\mathbf{s}_i) + \varepsilon_i, \quad i = 1, \dots, n_{\epsilon}.$$

- s: statistics vector
- $\varphi(\cdot)$ : conditional expectation  $\mathbb{E}[\theta|\mathbf{s}]$
- ε<sub>i</sub>: the residuals
- Adjusted samples:  $\tilde{\theta}_i = \hat{\varphi}(\mathbf{s}_{obs}) + \hat{\varepsilon}_i$



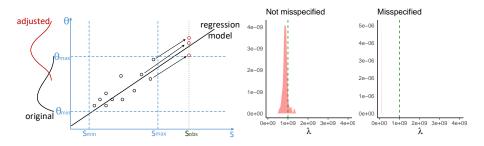
Model misspecification

- Regression-ABC can yield erroneous results under misspecification.
- Approx. posterior can lie beyond prior range.



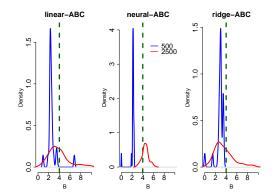
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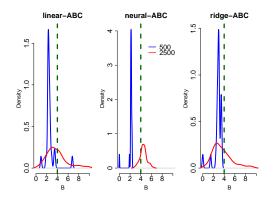
Low-simulation regime

- When model is expensive, number of simulations are limited.
- Regression is susceptible to overfitting.
- ABC posteriors can be concentrated far from the true value.



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Can we mitigate these issues by leveraging domain expertise?

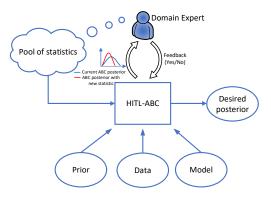
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### Human-in-the-loop (HITL) ABC

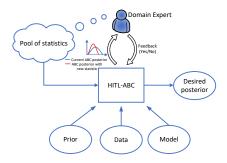
- Involve the expert in the inference procedure
- Elicit domain knowledge about statistics
- Assumptions:
  - Expert knowledge is tacit
  - Querying the expert is costly



# Problem setting

Human-in-the-loop (HITL) ABC

- Pool of statistics:  $S = \{s_1, s_2, \dots, s_w\}$
- Indicator of inclusion or exclusion of  $s_j$ :  $\gamma_j \in \{0, 1\}$
- With each vector **s**, we have associated  $\boldsymbol{\gamma} = [\gamma_1, \dots, \gamma_w]^\top$
- ABC posterior:  $p^{\epsilon}_{\mathsf{ABC}}( heta|\mathbf{y}_{\mathrm{obs}}, m{\gamma})$
- Let  $\gamma^*$  represent the "desired" statistics vector
- Goal: converge towards  $\gamma^*$  by querying expert about  ${\mathcal S}$



### Expert feedback model

Human-in-the-loop (HITL) ABC

Formulate expert feedback as a probabilistic model

- Expert provides binary feedback  $f_j \in \{0, 1\}$  about  $s_j$
- Model  $f_j$  as a noisy version of  $\gamma_j$ :

 $\gamma_j \sim \text{Bernoulli}(\rho_j),$  $f_j | \gamma_j \sim \gamma_j \text{Bernoulli}(\pi) + (1 - \gamma_j) \text{Bernoulli}(1 - \pi).$ 

- $\pi \in [0, 1]$ : the level of noise
- $\rho_j$ : prior probability of selecting  $s_j$
- Let  ${\mathcal F}$  be the set of feedback obtained
- $p(\gamma|\mathcal{F})$  available in closed-form

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ABC posterior based on  $\mathcal{F}$ :

$$p^{\epsilon}_{\mathsf{ABC}}( heta|\mathbf{y}_{\mathrm{obs}},\mathcal{F}) := \sum_{oldsymbol{\gamma} \in \{0,1\}^w} p^{\epsilon}_{\mathsf{ABC}}( heta|\mathbf{y}_{\mathrm{obs}},oldsymbol{\gamma}) p(oldsymbol{\gamma}|\mathcal{F})$$

#### Sequential experimental design

Human-in-the-loop (HITL) ABC

We design a sequential Bayesian experiment to select next statistic to query.

At iteration k + 1, we select  $s_{i^*}$ , where

$$j^* = \underset{j \notin \mathcal{J}_k}{\operatorname{arg\,max}} \quad \mathsf{KL}[p^{\epsilon}_{\mathsf{ABC}}(\theta | \mathbf{y}_{\mathrm{obs}}, \mathcal{F}_k, \tilde{\mathcal{F}}_j) \mid\mid p^{\epsilon}_{\mathsf{ABC}}(\theta | \mathbf{y}_{\mathrm{obs}}, \mathcal{F}_k)]$$

- $\mathcal{J}_k$ : set of indices of statistics queried after k iterations
- Estimate KL from samples as per [Perez-Cruz, 2008]

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- $\mathcal{J}_k$ : set of indices of statistics queried after k iterations
- Estimate KL from samples as per [Perez-Cruz, 2008]
- Stopping criterion: when utility of next statistic <  $\delta$
- Output: ABC posterior  $p^{\epsilon}_{\mathsf{ABC}}(\theta|\mathbf{y}_{\mathrm{obs}},\hat{\boldsymbol{\gamma}})$ , where

$$\hat{\gamma}_{k,j} = egin{cases} \arg\max p(\gamma_j|f_j), & ext{if } j \in \mathcal{J}_k \ \gamma_j \in \{0,1\} \ 0, & ext{otherwise}. \end{cases}$$

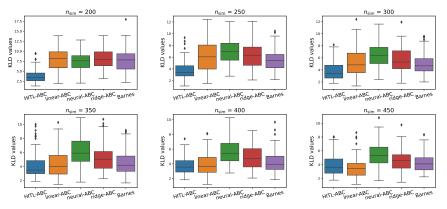
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Low-simulation regime

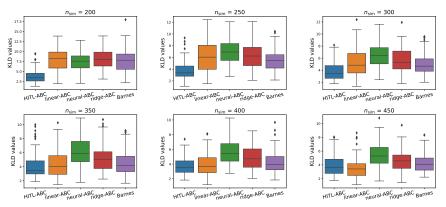
- Model: g-and-k distribution (4 parameters)
- Total 15 statistics in the pool (4 informative, 6 correlated, 5 noisy)
- HITL-ABC outperforms other methods for  $n_{\rm sim} \leq$  350, otherwise at par.
- Lack of simulations is compensated by expert's feedback.



#### Low-simulation regime

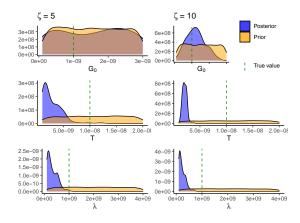
n <sub>sim</sub>	200	250	300	350	400	450
HITL-ABC	10.1	8.5	8.3	6.3	6.0	6.3
Random	13.8	13.6	13.4	13.3	13.1	13.4

#### Table: Average number of expert feedback required.

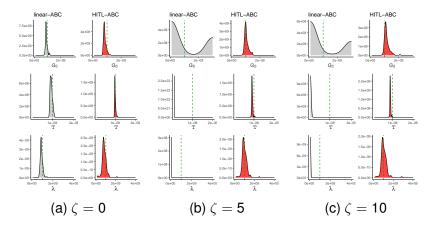


Model misspecification

- Radio propagation model (high-dimensional, complex-valued time series)
- Expert is shown inference results, and can detect misspecified statistics
- Parameters  $[G_0, T, \lambda]$ , 6 statistics (one mismatch)
- ζ: misspecification level



#### Model misspecification



When ζ > 0, performance of linear regression ABC seriously degrades

Removing mismatched statistic improves performance

#### Conclusion

- We introduce the first ABC method that actively leverages domain knowledge from experts in order to select summary statistics.
- With fairly limited effort from the expert (answering yes/no when presented with a few statistics), we are able to outperform the regression-ABC methods in situations where the simulation budget is low.
- Involving the experts in the ABC method gives us the opportunity to handle misspecified models, something the existing methods fail in.

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